Asymmetric competition causes multimodal size distributions in spatially-structured populations

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Abstract

Plant sizes within populations often exhibit multimodal distributions, even when all individuals are the same age and have experienced identical conditions. To establish the causes of this we created an individual-based model simulating the growth of trees in a spatially-explicit framework, which was parameterised using data from a long-term study of forest stands in New Zealand. First we demonstrate that asymmetric resource competition is a necessary condition for the formation of multimodal size distributions within cohorts. In contrast, the legacy of small-scale clustering during recruitment is transient and quickly overwhelmed by density-dependent mortality. Complex multi-layered size distributions are generated when established individuals are restricted in the spatial domain within which they can capture resources. The number of modes reveals the effective number of direct competitors, while the separation and spread of modes are influenced by distances among established individuals. Asymmetric competition within local neighbourhoods can therefore generate a range of complex size distributions within even-aged cohorts.

Keywords

bimodality; individual-based model; forests; Fuscospora cliffortioides; southern beech; zone-of-influence.
Introduction

Individual organisms within natural populations usually vary greatly in size. A description of the distribution of sizes is a common starting point for many demographic studies [e.g. 1, 2, 3]. This is especially the case for plants, where size distributions are often considered to convey information regarding the stage of development of a stand or the processes occurring within a population [4]. In the absence of asymmetric competition or size-related mortality, the sizes of individuals within an even-aged cohort should be approximately normally-distributed around a single mode, allowing for some variation in growth rate. More commonly a left-skew is observed during early stages of cohort development; this is attributed to smaller-sized individuals receiving insufficient resources to maintain growth, ultimately increasing their likelihood of mortality [5, 6]. Size-thinning thereafter reduces the degree of skewness [7, 8, 9] such that the distribution converges on a common maximum size [2]. Finally, as individuals die through disturbance or senescence, and recruitment into lower size classes occurs, populations shift to a size distribution referred to as reverse J-shaped, where a high density of of small individuals is combined with a small number of large dominants. This is a common pattern in old-growth forests, especially those dominated by shade-tolerant species which can persist in small size classes [e.g. 10].

A range of statistical models exist to capture these transitions in size distributions [4, 11]. Nevertheless, such simple models are unable to capture the behaviour of many systems. Multimodality of size distributions is widely observed in nature [2, 7, 12]. This is particularly true of plant populations [see Table 1 in 13], even when all individuals are known to have recruited simultaneously [14]. The preva-
ence of multimodality is likely to have been underestimated due to a failure to apply appropriate statistical tests [e.g. 15]. In some studies, even when multimodal distributions are observed, they are overlooked or dismissed as anomalous [e.g. 8, 11, 16].

When larger organisms monopolise access to resources it increases the asymmetry of competition among individuals [17]. Small individuals face combined competition from all neighbours larger than themselves, whereas large individuals are unaffected by their smaller neighbours. This is particularly likely to be the case for light competition among vascular plants, where taller stems capture a greater proportion of available radiation and determine access for those beneath [18]. As larger individuals can thereby maintain higher growth rates, incipient bimodality will be reinforced [12], at least until light deprivation causes mortality among smaller individuals [1]. Stand development models are able to generate bimodal patterns when resources for growth become limited [19, 20, 21]. Nevertheless, though the potential for bimodality to arise from competitive interactions is well-known, previous models have only been able to reproduce it within a narrow range of parameters [19, 20], leading to the conclusion that it is the least likely cause of bimodality in natural size distributions [12]. A range of alternative mechanisms might give rise to multimodality, including abiotic heterogeneity whereby large stem sizes are indicative of favourable environmental conditions [22], or sequential recruitment of overlapping cohorts [12]. Finally, the initial spatial pattern of recruits may itself create complex variation in the sizes of individuals.

In this study we argue that instead of being unusual or aberrant, multimodality is an expected outcome whenever strong asymmetries in competition among individuals occur in cohorts of sessile species. We sought to determine the conditions
under which multimodal size distributions form in spatially-structured populations using an individual-based modelling approach. Such models have the potential to derive new insights into fundamental ecological processes as they often demonstrate emergent properties which cannot be predicted from population-level approaches [23]. In order to parameterise our models we used a long-term dataset of 250 plots in New Zealand in which the sizes of over 20,000 *Fuscospora cliffortioides* (Hook. f.) Heenan & Smissen (≡ *Nothofagus solandri* var. *cliffortioides* (Hook. f.) Poole) trees have been monitored since 1974 [9, 24, 25]. These data are used to obtain plausible parameters for our simulation model, which is then employed to explore the causes of multimodality in virtual populations.

Our predictions were that (a) the size distribution of individuals would carry a long-term signal of the spatial patterns at establishment, and that (b) asymmetries in competitive ability would increase the degree of bimodality, which once established would strengthen through time, until resource deprivation removed weaker competitors from the population. Finally, we aimed to test whether (c) manipulating the distance and number of competitors within local neighbourhoods would generate variation in the number and positions of modes within cohort size distributions. Through this work we demonstrate that complex size distributions with multiple modes can be generated within cohorts even in homogeneous environmental space and when individuals are initially arranged in a regular grid. We show that multimodality is not a transient phase, but is maintained for the projected lifespan of a cohort. Finally, we show that the eventual size reached by any individual depends upon interactions with others in its immediate neighbourhood throughout its lifetime.
Materials and methods

The simulation model

All parameters used in the text are summarised in Appendix 1. The growth model is derived from a basic energy conservation principle. We assume throughout that resources in the model refer to light (and therefore carbohydrates acquired through photosynthesis), though in principle the model could be extended to other resources with appropriate parameterisation. Recruitment and age-related senescence are not included in the model. The resources $E$ that an individual acquires in a unit of time $t$ are distributed between the resources used to increase its size $M_g$ and all other metabolic and maintenance costs $M_m$. This is expressed mathematically as a general energy budget $E = M_g + M_m$. Assuming that resource intake scales with biomass $m$ as $E_i \propto m^{3/4}$ [26], and a linear relation between maintenance costs and biomass $M_m \propto m$, we can write a simple individual growth rate equation:

$$\frac{dm}{dt} = am^{3/4} - bm$$

where $a$ and $b$ are constants and the units are chosen such that an increase of one unit in biomass requires one unit of resources. A mathematically equivalent model, but with slightly different interpretation, has been proposed previously [18, 27, 28]. Equation 1 describes the potential growth rate of an individual in the absence of competition.

The potential rate of energy uptake of an individual is reduced when it competes with neighbours and thus they share the available light. In order to take this into
account the growth rate in the presence of competition can be expressed as

\[
\frac{dm}{dt} = am^{3/4} - bm - \sum_j I(m, m_j, d_j)
\] (2)

where \(I_j\) represents the reduction in biomass growth of a given individual due to competition with another individual \(j\) of mass \(m_j\) and at a distance \(d_j\) from the focal tree. The competitive response is obtained by summing \(I_j\) over all interacting neighbours. We only took pairwise interactions into account, summed across all interactions for each individual. This maintained computational efficiency of the simulations [29]. An individual died if its maintenance needs \(M_m\) were not met, i.e. if \(am^{3/4} - \sum_j I(m, m_j, d_j) < bm\).

Spatially explicit interactions among individuals were modelled with a circular zone of influence (ZOI) where \(A\) represents the potential two-dimensional space within which a plant acquires resources in the absence of competition. Resource competition between an individual \(i\) and its neighbour \(j\) is defined as occurring when \(A_i\) overlaps with \(A_j\). Within the area of overlap, \(A^{(I)}\), resources are distributed among the two individuals, but not necessarily equally. A larger individual (greater \(m\)) will be a stronger competitor, for example by over-topping in light competition, but also potentially through directing greater investment into below-ground resource capture [e.g. 30]. To incorporate asymmetric competition we define \(f_m(m, m_j)\) as being the proportion of resources \(E\) that an individual of size \(m\) obtains from the area within which it interacts with another individual of size \(m_j\). Assuming homogeneous resource intake within \(A\), then \(E\) is proportional to \(A^{(o)} + f_m(m, m_j)A^{(I)}\), where \(A^{(o)}\) is the area within which no interaction occurs \((A - A^{(I)})\).
Since in the absence of competition $E = am^{3/4}$, competition will reduce $E$ as follows:

$$E = am^{3/4} - (1 - f_m(m, m_j))A^{(f)}$$

and

$$I(m, m_j, d_j) = (1 - f_m(m_j))A^{(f)}$$

The explicit functional form for asymmetric competition is $f_m(m, m_j) = \frac{m^p}{m^p + m_j^p}$.

When $p = 0$ the resources in the zone of overlap are divided equally, irrespective of each individual’s size. If $p = 1$ then each individual receives resources in proportion to its size, and if $p > 1$ then larger individuals gain a disproportionate benefit given their size. This differs from a previous formulation [31], though their terminology of competitive interactions can be matched to this work as absolute symmetry ($p = 0$), relative symmetry ($p = 1$) and true asymmetry ($p > 1$). The shape of the competition kernel is identical in all cases.

This mathematical framework was used to create a spatially-explicit simulation model in which the growth and interactions among large numbers of individuals could be assessed simultaneously.

**Model fitting**

To obtain realistic parameters for the simulation model we utilised data from monospecific *Fuscospora cliffortioides* forests on the eastern slopes of the Southern Alps, New Zealand. *F. cliffortioides* is a light-demanding species which recruits as cohorts in large canopy gaps, and has a lifespan that seldom exceeds 200 years. The
data consisted of records from 20,330 trees situated in 250 permanently marked plots that randomly sample 9,000 ha of forests. Each plot was 20×20 m in size. In the austral summer of 1974–75 all stems >3 cm diameter at breast height (dbh) were tagged and dbh recorded. The plots were recensused during the austral summers of 1983–84 and 1993–94. Only stems present in more than one census were included. Previous work on this system has confirmed a dominant role for light competition in forest dynamics [9, 18].

We tested the tree size distribution from the first survey of each plot for multimodality by fitting a finite mixture model of one, two and three normal distributions (see Appendix 1). We employed an expectation-maximisation (EM) algorithm [32] within the R package FlexMix 2.3-4 [33] and utilised the Bayesian Information Criterion (BIC) to decide whether each size distribution was unimodally or multimodally distributed.

In order to fit the simulation model to the data we estimated the mass $m$ of the trees by allometric relation $\text{dbh} = C_{\text{dbh}}m^{3/8}$ [26, 34], where $C_{\text{dbh}}$ was taken as a free parameter. The area $A$ of the circle representing the potential space for resource acquisition was given by $cA = am^{3/4}$ where $c$ is a proportionality constant. A linear relation between dbh and radius of the zone of influence was chosen, and a high degree of asymmetric competition ($p = 10$). The latter improved overall fit of the models, indicating a role for asymmetric competition in driving stand dynamics.

For each of 250 plots we began the simulation model with the observed stem sizes from 1974 attached to points randomly distributed in space. The simulation was run for 19 model years, developed in time increments $\delta t$ which nominally correspond to 10 weeks (for simplicity there is no seasonality of growth). An
individual’s growth is given by:

\[
\delta m_i = \left[ am_i^{3/4} - bm_i - \sum_j \frac{m_j^p}{m_i^p + m_j^p}cA_j^{(I)} \right] \delta t
\]  

(5)

In each Monte Carlo iteration individuals \( m_i \) were selected at random and their size updated. A search algorithm was employed to find values of \( a \) and \( b \) which gave the best fit to the observed individual growth rates with Pearson’s \( \chi^2 \), averaged across the ensemble of simulations. Note that the model was fit to the growth rates of individual stems based on repeated measurements, rather than stand-level properties such as size distributions.

Having obtained suitable values for \( a \) and \( b \) we performed simulations to compare the size distributions as predicted by the model (assuming initially random stem positions) with the empirical distributions observed in the data set. These were initiated using size distributions from stands in the \( F. 
cliffortioides \) dataset in which the mean stem diameter was small (\( \bar{d} < 15 \) cm), then run until the mean reached a medium (\( 15 \) cm \( \leq \bar{d} < 22 \) cm) or large (\( \bar{d} \geq 22 \) cm) stem size. Estimates of size-dependent mortality rate were also obtained and compared with empirical outputs as in [9]. This provides an independent evaluation of model performance as mortality rates were not used to parameterise the model.

Exploring multimodality in size structure

The simulator with fitted parameters as described above was used to explore the factors which cause multimodal size distributions to form. We tracked the development of size structures in simulated stands with differing initial spatial patterns and symmetry of competition. In these simulations all individuals were of identical
First 2100 spatial patterns were generated, each containing a distribution of
points with $x$ and $y$ co-ordinates in a virtual plot of $20 \times 20$ m. Equal numbers
patterns were clustered, random and dispersed. Random patterns were produced
using a uniform Poisson process with intensity $\lambda = 0.05$ points $m^{-2}$. Clustered
patterns were created using the Thomas process. This generated a uniform Poisson
point process of cluster centres with intensity $\lambda = 0.005$. Each parent point was
then replaced by a random cluster of points, the number of points per cluster being
Poisson-distributed with a mean of 10, and their positions as isotropic Gaussian
displacements within $\sigma = 1$ from the cluster centre. Dispersed patterns were
produced using the Matern Model II inhibition process. First a uniform Poisson
point process of initial points was generated with intensity $\lambda = 0.06$. Each initial
point was randomly assigned a number uniformly distributed in $[0,1]$ representing
an arrival time. The pattern was then thinned by deletion of any point which
lay within a radius of 1.5 units of another point with an earlier arrival time.
All patterns were generated in $\mathbb{R}$ using the spatstat package [35]. Each pattern
contained roughly 500 points (clustered $N = 501 \pm 2.7$, random $N = 501 \pm 0.8$,
dispersed $N = 488 \pm 0.7$). The slightly lower number of points in the dispersed
pattern reflects the inherent difficulties in generating a dense pattern with a highly-
dispersed structure and has no qualitative effect on later analyses. Although the
density within starting patterns was approximately a quarter of that observed in
the empirical data, initial density has a limited effect on final outcomes since its
main effect is to reduce the time until points begin to interact [36], and lower point
densities increased computational speed, allowing for greater replication.

A number of further patterns were generated to explore the influence of specific
parameters. First, a regular square grid was used with a fixed distance of 1.5 or 3 m between individuals. Next, groups of individuals were created in which all individuals within groups were 3 m apart, but with sufficient distance among groups that no cross-group interactions could take place. Groups contained either two individuals (pairs), three individuals in a triangular arrangement (triads) or four individuals in a square arrangement (tetrads). The total starting population in each pattern was approximately 7500 individuals.

We ran simulations of the spatially-explicit individual-based model, varying the degree of asymmetric competition $p$. The points generated above became individual trees represented as circles growing in two-dimensional space. Each individual was characterised by its mass $m$ and co-ordinates. In order to model mortality, an individual was removed from the simulation if carbon losses exceeded gains, that is if $\left[ am_i^{3/4} - bm_i - \sum_j \frac{m_i^p}{m_i^p + m_j^p} A_{ij}^I \right] < 0$.

The predicted size distribution and mortality rate of clumped, random and dispersed starting patterns were obtained from ensemble averages of 700 simulations corresponding to the point processes generated above. $m$ was a continuous variable but in order to derive the size distribution, growth and death rates we calculated size frequencies based on 10 kg biomass bins. Since the death rate changes through time due to alterations in the size structure of the community, we present the average death rate for each size class across all time steps in simulations, which run for 460 model years (at which point only a few very large stems remain). This allows sufficient resolution for figures to be presented as effectively continuous responses rather than histograms, and is equivalent to a landscape-scale aggregation of size-dependent mortality data across a series of stands of differing ages.
Results

Analysis of the New Zealand forest plot dataset revealed multimodal distributions in 179 plots in 1974, 163 plots in 1984 and 152 plots in 1993 from of a total of 250 plots in each survey. This represents 66% of plots, showing that multimodality is more common than unimodality within these forests (see Appendix 2).

The simulation model was fit to the observed individual growth rates in the *F. cliffortioides* dataset and provided a robust representation of the empirically-measured patterns. The fitted parameters (*a*, *b* and *Cdbh*) are shown in Appendix 1. The effectiveness of the model was assessed through its ability to capture size-dependent mortality rates, which were an emergent property of the system and not part of the fitting process. Size distributions thus obtained were qualitatively similar to those observed in the empirical dataset [9]; see Appendix 3.

Subsequent simulation modelling used the parameters derived from the *F. cliffortioides* dataset (*a*, *b*, *Cdbh*) and created simulated forests to investigate the potential origins of multimodal patterns. Using stochastically-generated starting patterns, major differences were evident in the patterns of growth and survival depending on the degree of competitive asymmetry *p* and the initial spatial configuration (Fig. 1).

With completely symmetric competition among individuals (*p* = 0), average tree growth in clustered patterns was greater than in either random or dispersed patterns (Fig. 1a). This unexpected result can be attributed to the high rate of density-dependent mortality in very early time steps (Fig. 1d). Initial mortality in random patterns reduced the population to be comparable with dispersed patterns, compensating for the slight initial differences in abundance. Clustered populations
remained larger in average stem size (Fig. 1a) as the result of a smaller final population size (Fig. 1d), an effect which developed rapidly and was maintained beyond the plausible 200-year lifespan of *F. cliffortioides*.

In the absence of asymmetric competition (*p* = 0), starting patterns had a limited effect on final size distributions, with only minor increases in skewness in clustered populations at advanced stages of development (Appendix 4). In all cases size distributions remained unimodal. It is therefore apparent that variation in initial spatial patterns is not in itself sufficient to generate multimodality in size distributions, at least not unless the average distance among individuals exceeds their range of interaction, which is highly unlikely in the context of plant populations.

The introduction of weak asymmetry (*p* = 1) tended to increase the mean size of individuals while causing reductions in population size (Fig. 1b,e) and diminishing the differences among initial patterns, such that with strong asymmetry (*p* = 10) the differences in final size between starting patterns were negligible (Fig. 1c). Strong asymmetry also caused population sizes to converge within the likely lifespan of the trees, irrespective of starting conditions, and at a lower final level (Fig. 1f). Reduced differences among initial patterns with increasing asymmetry arose because fewer small trees survived around the largest tree in the vicinity, which caused patterns to converge on a state with dispersed large individuals and smaller individuals in the interstices. More left-skewed distributions also emerged as a consequence of the low tolerance of individuals to depletion of resources (individuals failing to obtain sufficient resources for their metabolic needs died immediately). Thus the small individuals die soon after their resource acquisition area is covered by the interaction range of a larger individual. Such
left skew would be reduced for species capable of surviving long periods of time with low resources either through tolerance or energy reserves.

Increasing competitive asymmetries caused size distributions to exhibit slight multimodality with a lower frequency of individuals in the smaller size class at 150 years (Appendix 5). Given entirely random starting patterns, more pronounced bimodality emerged as the degree of asymmetric competition increased. Furthermore, the model predicted a U-shaped size-dependent mortality rate, qualitatively consistent with a pattern in the empirical data (Appendix 6; compare Fig. 5 in [9]). This trend intensified with increasing asymmetric competition, and was absent when resource division was symmetric. It occurred because in large trees the majority of resources are required for maintenance, and therefore even a relatively small amount of competition ultimately increases their mortality rate. In the absence of asymmetric competition the death rate of large trees approached zero.

Greater insights into the causes of multimodality are revealed through the use of designed spatial patterns in which the timing of interactions within model development can be precisely controlled. These illustrate that the separation between modes is determined by the distance among competing individuals under asymmetric competition (Fig. 2 and Appendix 7). The size structure can therefore provide an indication of the dominant distance over which individuals are competing, though separation of modes will be less clear when a strict grid is absent. Note that the position of the right-hand mode remains identical, and it is only the mode of the subordinate individuals which shifts to a smaller size class. Highly-dispersed patterns give rise to more complex size distributions through their development when asymmetric competition is present. In the most extreme case, when initial
patterns are gridded, each individual interacts with a series of neighbours as its size increases, leading to a complex multimodal pattern, at least until continued mortality removes smaller size classes (Fig. 3). Note that the modes are more clearly distinguished than is the case for random starting patterns where distances among individuals vary (compare Fig. 8c).

The patterns generated by small groups of interacting individuals at equal distances apart with asymmetric competition lead to size distributions with a number of modes equal to the number of individuals within each group. For patterns derived from pairs of individuals, the size distribution is bimodal, and in similar fashion triads and tetrads produce size distributions with three and four modes respectively (Fig. 4). Each mode corresponds to the discrete ranking of individuals within groups. This indicates that in gridded populations, as might be observed in plantations or designed experiments, the number of modes is determined by the effective number of competitors.

Discussion

Multimodality in cohort size distributions is the outcome, rather than the cause, of asymmetric competition among individuals of varying size. Regardless of initial small-scale starting patterns, size distributions remain unimodal in the case of symmetric competition among individuals. Only when larger individuals are able to acquire a greater proportion of resources from shared space does bimodality begin to emerge. Spatial patterns of established individuals can modulate these interactions, with complex multimodal distributions generated when individuals are either regularly or highly dispersed in space. The number of modes
corresponds to the number of effective competitors and their separation is a consequence of average distances among individuals. Note that our simulations do not incorporate continuous recruitment; this is a reasonable assumption for systems such as *F. cliffortioides* forests, where large-scale disturbances are followed by stand replacement.

Asymmetric competition will lead to multimodal distributions at some point during stand development. We extend upon previous studies [e.g. 37] by providing a general framework for predicting and interpreting complex size distributions in spatially-structured and even-aged populations. Under light competition the modes will correspond to discrete and well-defined canopy layers. In [13] a series of controlled experiments were conducted to investigate size distributions in populations of annual plants, finding in many cases that distributions with two or three modes were observed. Our results allow for a fuller interpretation of these earlier findings, as we have shown that the number of modes reflects the number of effective competitors, placing a limit on the complexity of size distributions.

As demonstrated in Figs. 2 and 4, the larger mode remains in the same position regardless of the size at which competition begins. This highlights that those individuals in larger size classes are almost unaffected by competition during stand development.

Even when all individuals in a cohort begin with identical size, small fluctuations in the acquisition of shared resources lead to a multimodal size distribution, regardless of whether the initial pattern was random, dispersed or clustered. The size distribution is not affected by differences in the initial spatial structure at small scales due to the death of close neighbours early in stand development. A similar result was found by [36], who argue that the importance of recruitment patterns in
generating asymmetries in competition may have been over-stated. Likewise initial
density will have a limited effect on final size distributions as its main influence is
on the time at which individuals begin to interact [36]. Therefore, while local in-
teractions undoubtedly do cause competitive asymmetries [e.g. 17], these are more
relevant in determining the pattern of mortality during self-thinning rather than
final size distributions, so long as the distances over which competition influences
growth are larger than the characteristic scales at which initial spatial structuring
occurs. In dense aggregations of recruiting plants this is likely to be the case.

The model implies only a single resource for which individuals compete. It is
typically assumed that above-ground competition for light is asymmetric, whereas
below-ground resources are competed for symmetrically [38], though the latter
assumption may not always be true [e.g. 39, 40]. More complex zone-of-influence
models can take into account multiple resources and adaptive allometric changes on
the part of plants in response to resource conditions [e.g. 41, 42]. Indeed, plasticity
can diminish the impact of asymmetric competition [41, 43]. Although below-
ground interactions are challenging to measure directly, there is good evidence
that above- and below-ground biomass scale isometrically [44] which justifies the
use of above-ground biomass to infer potential root competition. Previous work
using the same data has identified a dominant role for light competition among
smaller stems, with nutrient competition important at all stem sizes [18].

Forest mensuration tends to overlook the shape of size distributions in favour
of summary statistics [e.g. mean size, coefficient of variation, maximum size; 45]
and may therefore miss out on valuable contextual information. While the utility
of size distributions as a predictive tool for modelling dynamics has been fre-
quently overstated [see 46], they can nonetheless remain a valuable indicator of
past dynamics. One outcome of bimodality arising from asymmetric competition is that large and small individuals have differing spatial patterns, with the larger dispersed in space and the smaller confined to the interstices generated by the dominant competitors [47]. This can be used as a diagnostic tool as it allows this mechanism to be distinguished from abiotic heterogeneity, leading to clustering of similar sizes, or independent sequential recruitment, leading to a lack of co-associations between size classes [12]. Likewise in mixed-species stands succession can cause a multimodal pattern to emerge through aggregation of several unimodal cohorts, persisting throughout stand development [10]. The interplay between size distributions, plant traits and disturbance can generate complex emergent patterns in forest dynamics at the landscape scale [48]. Bimodality generated by size competition among individuals is a distinct phenomenon from the bimodality in inherited size across species which is often observed in mixed-species communities [e.g. 49]. Where size histograms combine individuals from multiple species, the causes of bimodality are likely to include long-term evolutionary dynamics in addition to direct competition among individuals. Contextual information on spatial patterns, disturbance regimes and community composition are therefore essential to interpreting size distributions in natural systems.

Our models are based upon parameters obtained from a long-term dataset and can therefore be immediately transferred to a predictive framework. While the exact terms are most suited to the Fuscospora cliffortioides forests which form the basis of this work, it is likely that they will be applicable to any monospecific plant population. Bimodal size distributions might be overlooked where aggregate curves are drawn as composites of a large number of plots, which will tend to average out differences, or where appropriate statistical tests are not employed.
We find that 66% of plot size distributions in our data are bimodal. It is likely that these do not all represent single cohorts; for example, a severe storm in 1972 opened the canopy in some plots and allowed a recruitment pulse [24, 50]. Irrespective of this, our growth model is able to capture subsequent stand development regardless of the origin of the bimodality (see Appendix 3). Our results also show that multimodality can act as an indicator of asymmetric competition. Thomas & Weiner [31] present evidence that the degree of asymmetry in natural plant populations is strong, with larger individuals receiving a disproportionate share of the resources for which they compete ($p \gg 1$). The phenomenon of multimodality should therefore be widespread.

In conclusion, and in contrast with a previous review of bimodality in cohort size distributions [12], we contend that asymmetric competition is the leading candidate for explaining multimodal size distributions, and is its cause rather than the outcome. Previous simulation results suggesting that the parameter space within which multimodality occurs is limited were based on stand-level models. Through the use of individual-based models it can be demonstrated that multimodality is an expected outcome for any system in which larger individuals are able to control access to resources, and where individuals compete in space. The strength of these asymmetries determines the degree to which multimodality is exhibited, while the number and separation of modes are determined by the number of effectively-competing individuals and the distances among them. While multimodality may be a transient phase within the development of our models, many forest stands exhibit non-equilibrial conditions, and indeed most natural plant populations are prevented by intermittent disturbance from advancing beyond this stage [24, 50]. Consistently unimodal size distributions should be seen as the exception rather
than the rule.

**Data accessibility**

Data were obtained from New Zealand’s National Vegetation Survey Databank and can be accessed at http://datastore.landcareresearch.co.nz/dataset/multimodal-size-distributions-in-spatially-structured-populations. All C code used to run the simulations can be obtained from https://github.com/jorgevc/IMB-SizeDependent.

**Competing interests**

We have no competing interests.

**Authors’ contributions**

MPE and JV conceived and designed the study; RBA and DAC provided data; JV carried out the statistical analyses under guidance from MPE, DAC and RBA; JV and MPE prepared the first draft of the manuscript. All authors contributed towards manuscript revisions and gave final approval for publication.

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Figure captions

**Figure 1.** Cohort-level characteristics of stands with either random, clustered or dispersed initial starting patterns over $t$ years (simulation time). (a–c) Mean tree size in kg with increasing levels of asymmetric competition $p$ (0, 1, 10), note that (a) has a reduced $y$-axis length; (d–f) mean number of surviving individuals $N$ per $20 \times 20$ m plot with competition varying from symmetric ($p = 0$) to weakly ($p = 1$) and strongly asymmetric ($p = 10$). Each line is derived from an ensemble average of 700 simulations.

**Figure 2.** Separation between modes with varying distance of competing neighbours and strong asymmetric competition ($p = 10$). Size distributions of stands composed by pairs of equidistant individuals after 200 years of development. Solid line: individuals spaced at 1.5 m, dashed line: individuals spaced at 3 m. Each line is derived from an ensemble average of 700 simulations.

**Figure 3.** Emergent size distribution through stand development given an initially gridded starting pattern. Individuals separated by 1.5 m from their neighbors and with strong asymmetric competition ($p = 10$). Panels show distribution at 150, 200, 230 and 250 years. Each plot is derived from an ensemble average of 700 Monte Carlo simulations.

**Figure 4.** Size distributions of stands composed of groups of two, three and four equidistant competing individuals (pairs, triads and tetrads respectively) with 3 m of separation among individuals in each group. Asymmetric competition set
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## Appendix 1

Table 1: Model terms as used in the text, separated between fitted parameters obtained from field data and free variables at the individual and stand level.

<table>
<thead>
<tr>
<th>Symbol</th>
<th>Value</th>
<th>Units</th>
<th>Definition</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Fitted parameters</td>
<td></td>
<td></td>
</tr>
<tr>
<td>$a$</td>
<td>$2.5 \times 10^{-3}$</td>
<td>$10\times \text{kg}^{-3/4}\text{year}^{-1}$</td>
<td>Conversion factor between $m^{3/4}$ and $E$</td>
</tr>
<tr>
<td>$b$</td>
<td>$2.5 \times 10^{-4}$</td>
<td>$10\times \text{kg}^{-1}$</td>
<td>Resource cost for maintenance per unit biomass</td>
</tr>
<tr>
<td>$C_{dbh}$</td>
<td>9.4</td>
<td>$\text{cm}/10\times \text{kg}^{3/8}$</td>
<td>Allometric relation between biomass and dbh</td>
</tr>
<tr>
<td></td>
<td>Individual-level parameters</td>
<td></td>
<td></td>
</tr>
<tr>
<td>$m$</td>
<td>variable</td>
<td>$10\times \text{kg}$</td>
<td>Biomass of an individual</td>
</tr>
<tr>
<td>$d_{ij}$</td>
<td>variable</td>
<td>m</td>
<td>Distance of an individual $i$ to its neighbour $j$</td>
</tr>
<tr>
<td>$A_{ij}$</td>
<td></td>
<td>$\text{m}^2$</td>
<td>Area of interaction between an individual $i$ and its neighbour $j$</td>
</tr>
<tr>
<td></td>
<td>Stand-level parameters</td>
<td></td>
<td></td>
</tr>
<tr>
<td>$p$</td>
<td>fixed</td>
<td>dimensionless</td>
<td>Degree of competitive asymmetry. $p = 0$ corresponds to symmetric competition while $p &gt; 0$ indicates asymmetric competition</td>
</tr>
<tr>
<td>$E$</td>
<td>equation (3)</td>
<td>$10\times \text{kg year}^{-1}$</td>
<td>Resource intake rate of an individual</td>
</tr>
<tr>
<td>$I(m, m_j, d_j)$</td>
<td>equation (4)</td>
<td>Resource year$^{-1}$</td>
<td>Reduction of resource intake rate due to competition</td>
</tr>
<tr>
<td>$f_m(m, m_j)$</td>
<td></td>
<td>dimensionless</td>
<td>Fraction of resources that an individual of biomass $m$ obtains from the area of interaction with an individual of biomass $m'$</td>
</tr>
</tbody>
</table>
Figure 5: Frequency of *Fuscospora clifortioides* plots in New Zealand exhibiting uni- or multimodality in size distribution as determined by finite mixture models testing for the presence of one, two or three modes. Data from initial 1974–1975 surveys.
Figure 6: Comparative figure to match Fig. 3 in Coomes & Allen 2007. Histograms show distributions of diameter at breast height (dbh; cm) of stands in which mean stem sizes were of medium (15–22 cm dbh; a) and large mean size (>20 cm dbh; b) in 1974. Simulations began with trees in random positions following a size distribution taken from the 117 stands with small mean stem size (<15 cm) in 1974. Dashed lines indicate patterns in simulated stands after 20 or 70 years of model time respectively. This is the ensemble average of $117 \times 4 = 468$ simulations.
Appendix 4

Figure 7: Mortality rate as a function of tree size. Solid line for symmetric competition, dashed and dotted lines correspond to increasing asymmetric competition. Derived from an ensemble average of 700 simulations, each of which is run for a nominal 460 years, and showing the cumulative function over the whole time period.
Appendix 5

Figure 8: Size-frequency histograms for simulated stands. All plots represent 150 years of stand development with increasing levels of asymmetric competition $p$ (0, 1, 10) and random initial pattern. Each plot is derived from an ensemble average of 700 simulations.
Figure 9: Size distributions of populations with symmetric competition among individuals ($p = 0$) but variation in initial pattern (random, dispersed, clustered). Panels show distribution at 150 and 200 years. Each plot is derived from an ensemble average of 700 simulations.
Figure 10: Effect of increasing distance between paired individuals within simulations (as Fig. 2) on separation between modes in the emergent size distribution. Note that increasing distance reduces the separation of modes by increasing the model time required for two individuals to begin competing for resources.